



FROM CHIPS TO SYSTEMS - LEARN TODAY, CREATE TOMORROW

DEC 5 - 9, 2021 🔶 San Francisco, California



Statheros: A Compiler for Efficient Low-Precision Probabilistic Programming

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ILLINOIS







Probabilistic Programs

• Extend normal programs with :



Probabilistic Programs

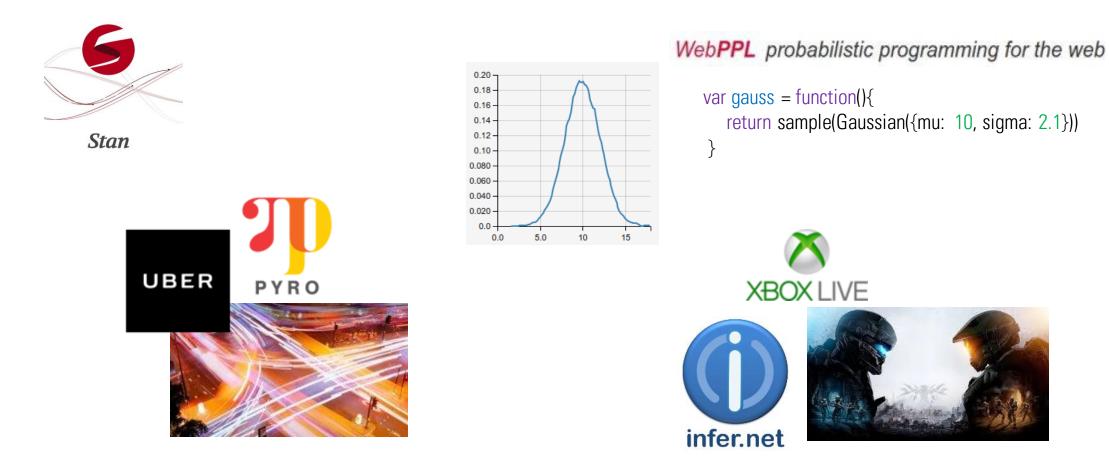
- Extend normal programs with :
 - Random Sampling x |= Normal(0,1);

Conditioning on Data Data<Fixed> y = {1.2,...}

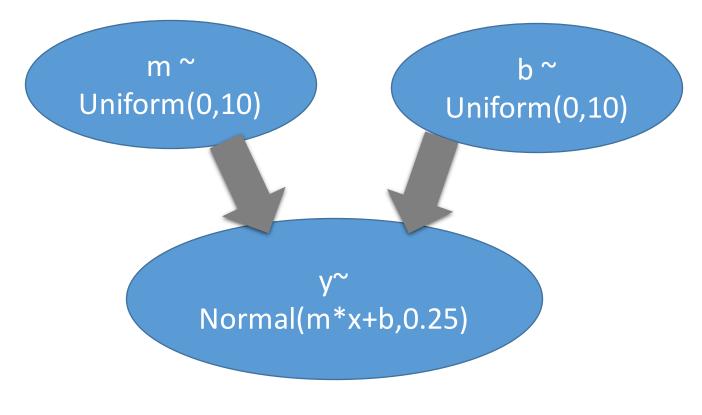
Posterior over Parameters Param<Fixed> x;



Probabilistic Programs



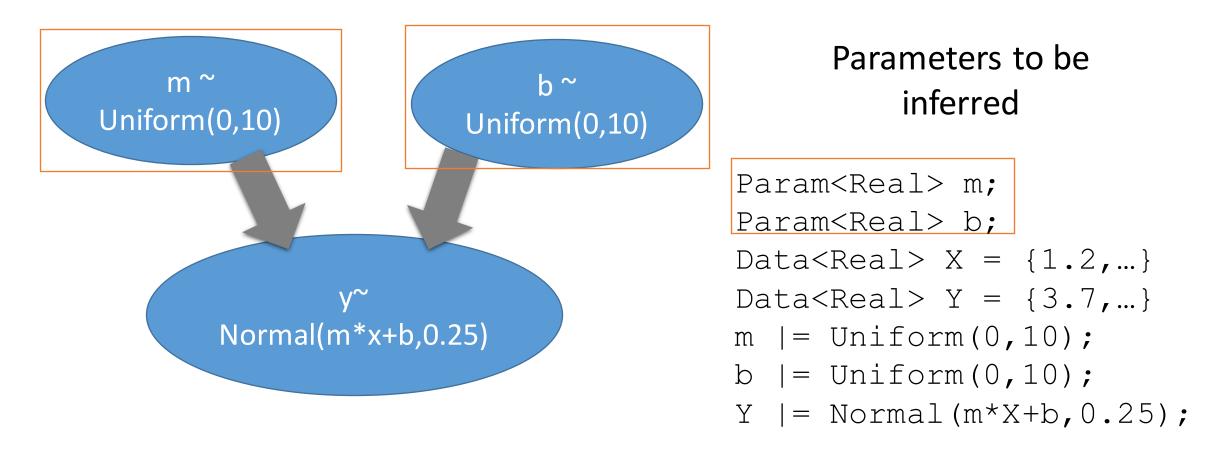




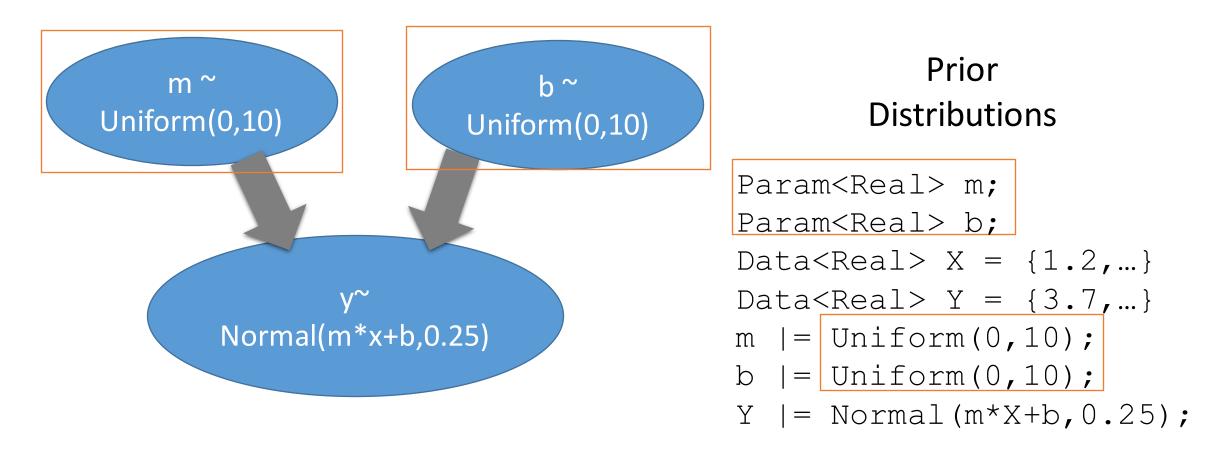
Param<Real> m; Param<Real> b; Data<Real> X = {1.2,...} Data<Real> Y = {3.7,...} m |= Uniform(0,10); b |= Uniform(0,10);

Y = Normal(m*X+b, 0.25);

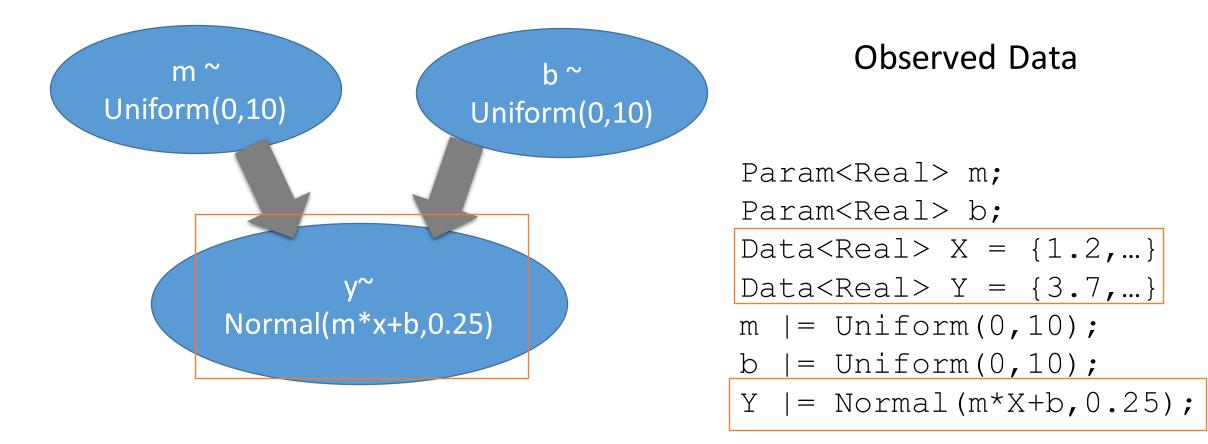




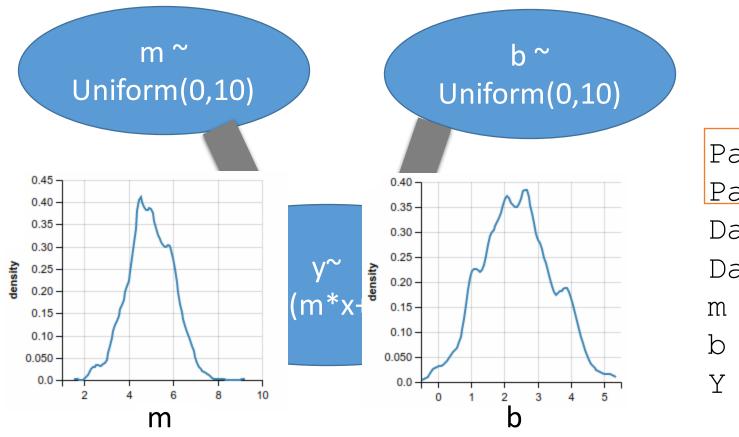












Posterior Distributions after Inference

Param<Real> m; Param<Real> b;

Data<Real> $X = \{1.2, ...\}$

Data<Real> $Y = \{3.7, ...\}$

- $m \mid = Uniform(0, 10);$
- b $\mid = \text{Uniform}(0, 10);$

Y = Normal(m*X+b, 0.25);



Edge Computing











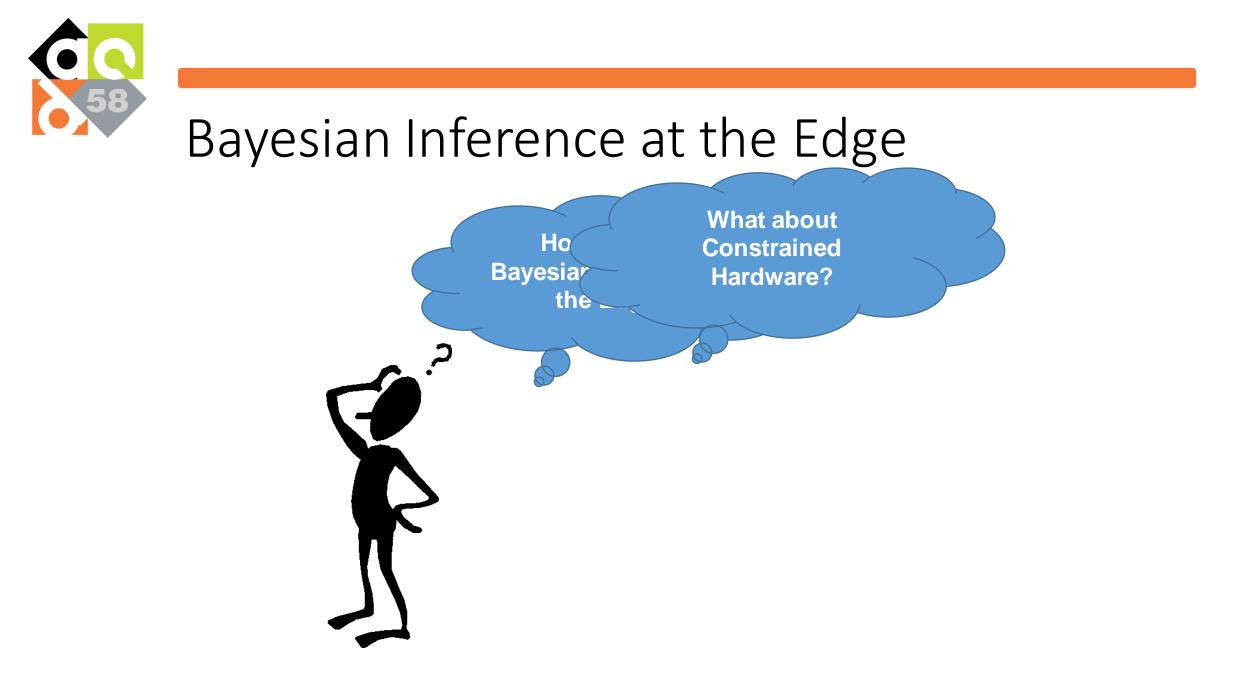


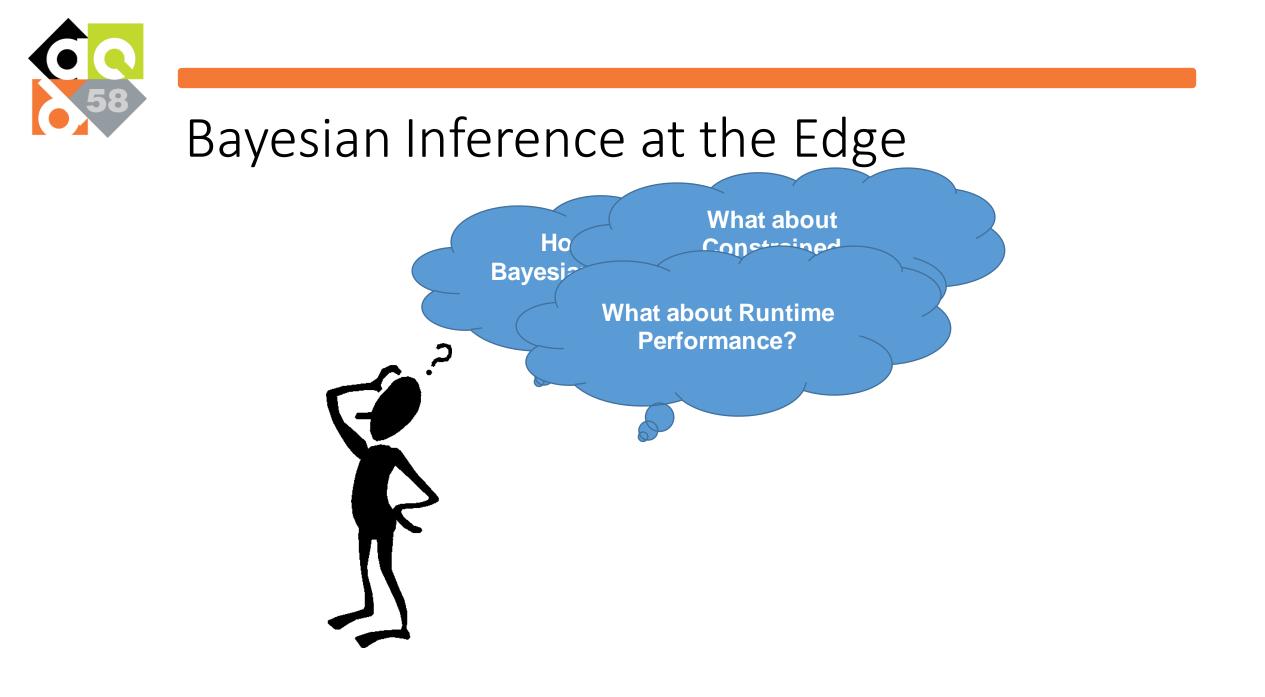
Bayesian Inference at the Edge

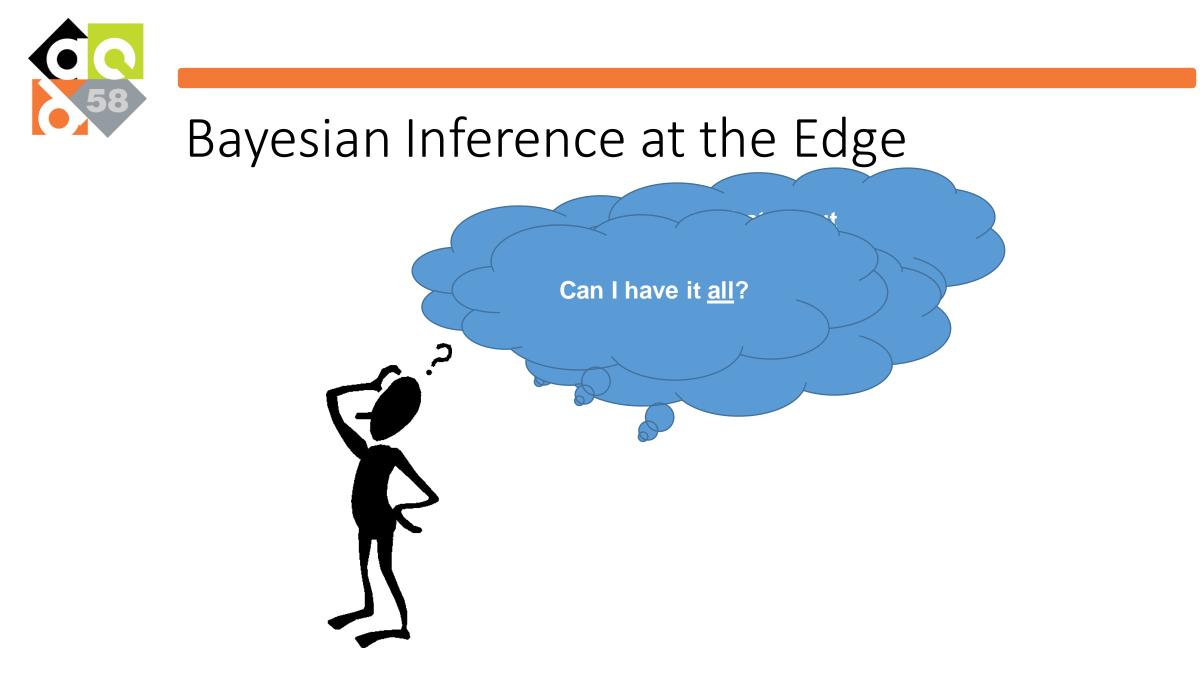














Idea 1)



Idea 1) Get a Ph.D. in ML + Embedded Systems



Idea 1) Get a Ph.D. in ML + Embedded Systems

Est. Time: 5-6 years



Idea 2)



Idea 2) Use fully automated Compiler framework



Statheros

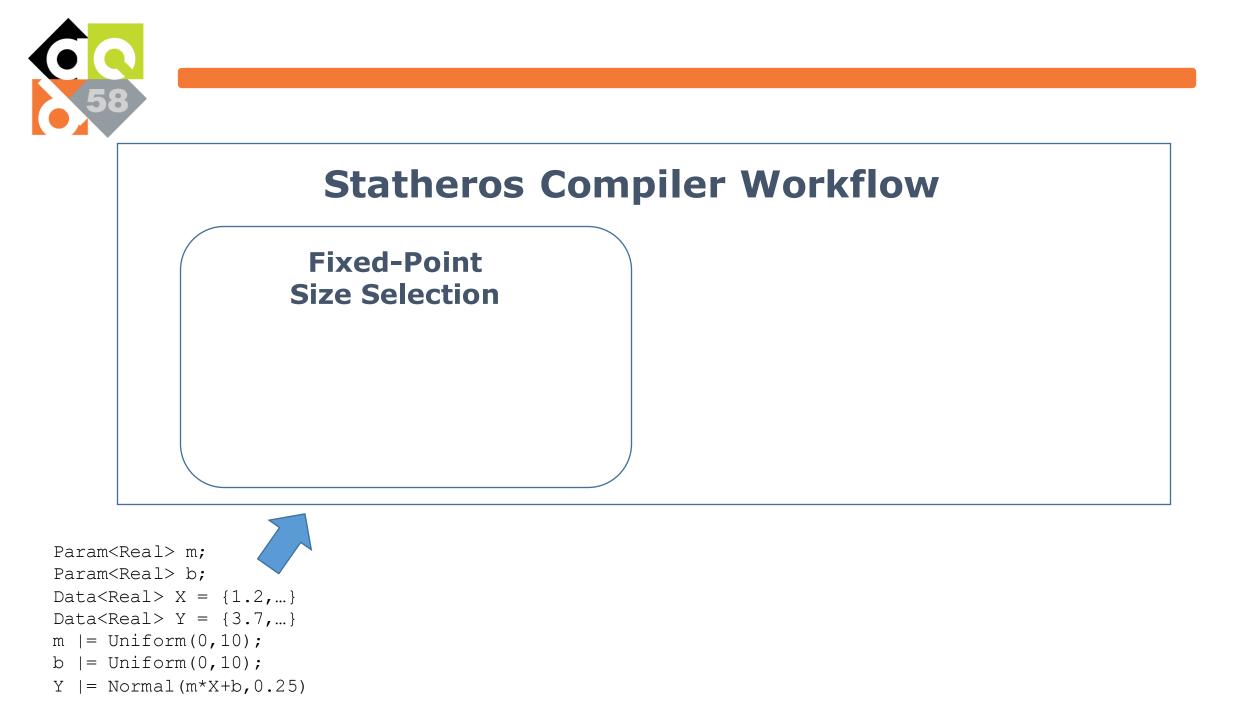
- Embedded in C++
- All MCMC code uses fixed-point arithmetic
- Optimal configurations inferred by compiler
- Full integration with ARM toolchain for edge-device processors

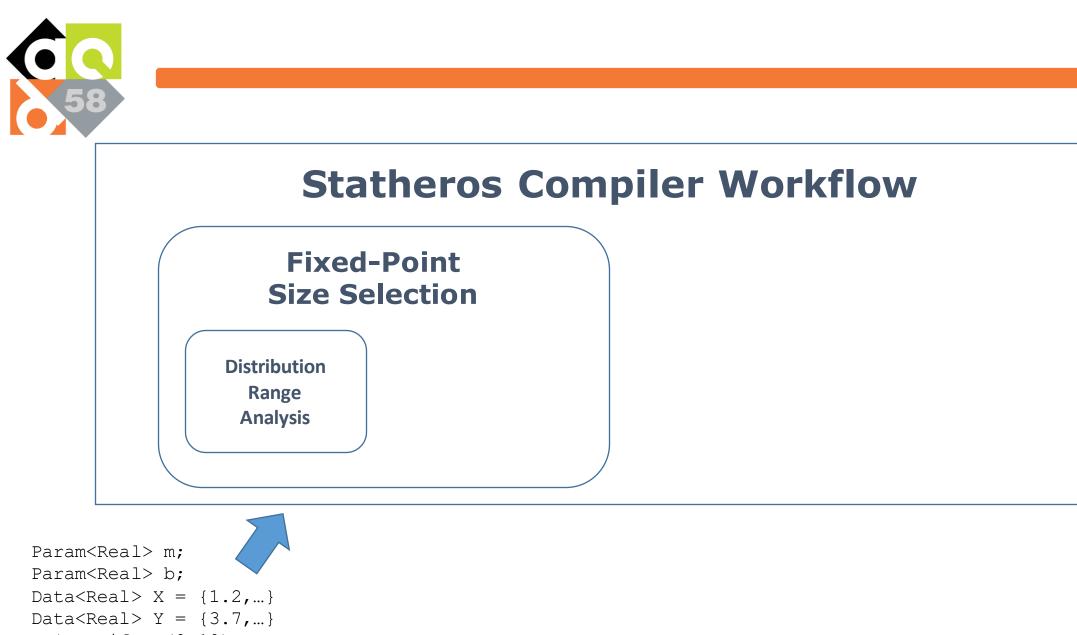


Statheros Compiler Workflow



Statheros Compiler Workflow

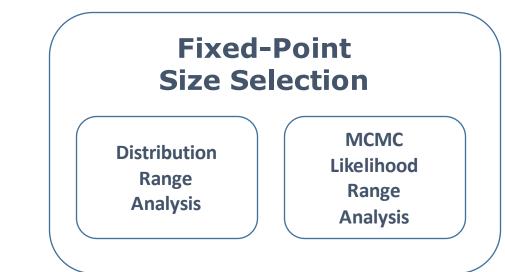




- m |= Uniform(0,10);
- b |= Uniform(0,10);
- $Y \mid = Normal(m*X+b, 0.25)$

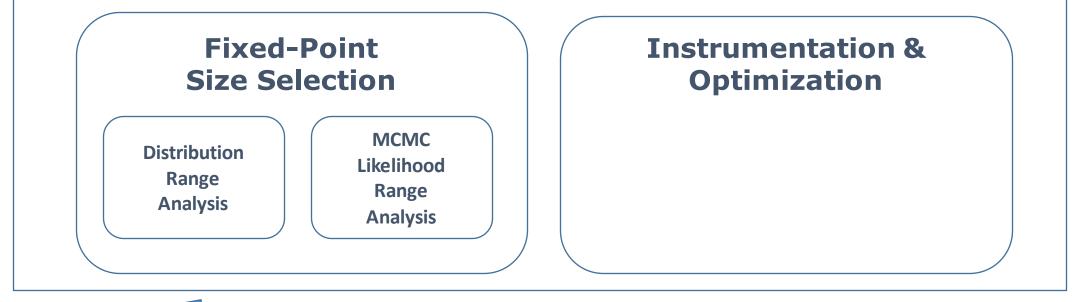






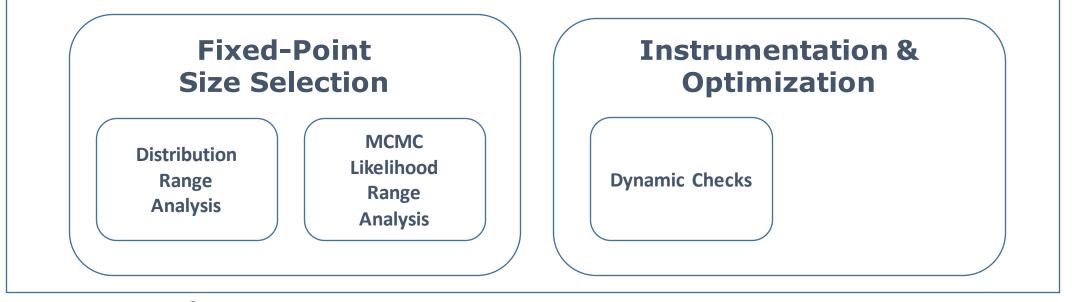






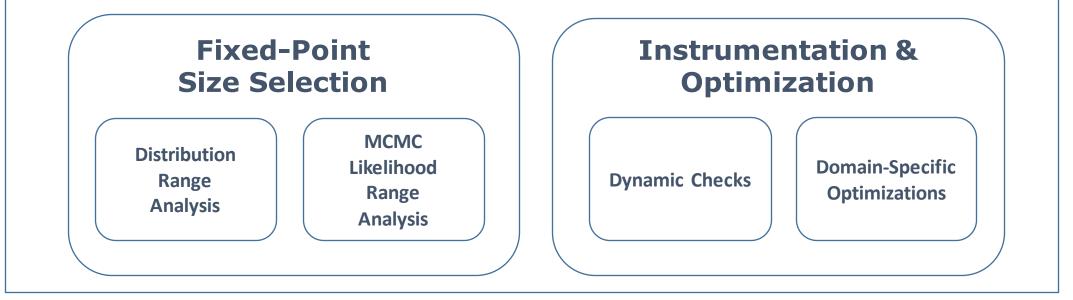






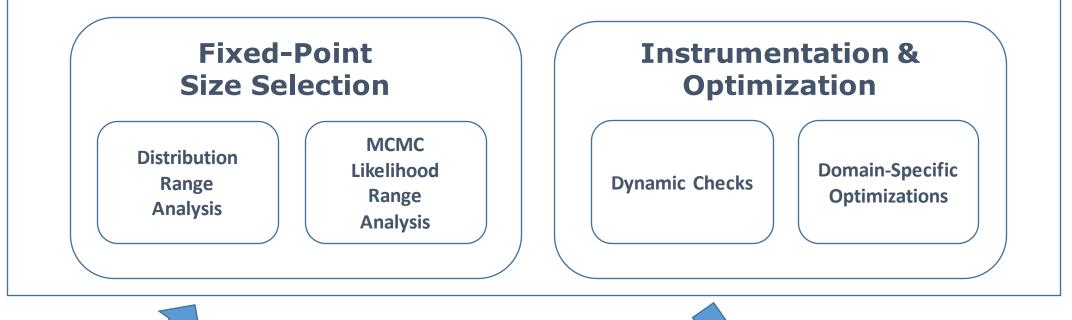


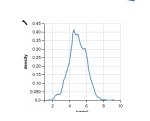












 0x57
 0x65
 0x20
 0x63
 0x6f
 0x6e
 0x76

 0x65
 0x72
 0x74
 0x20
 0x68
 0x69
 0x67

 0x68
 0x20
 0x6c
 0x65
 0x76
 0x65
 0x66

 0x20
 0x70
 0x72
 0x61
 0x62
 0x61
 0x62

 0x69
 0x70
 0x72
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 0x60
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 0x73
 0x74
 0x69
 0x63

 0x20
 0x70
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 0x64
 0x65

 0x64
 0x22
 0x20
 0x63
 0x66
 0x64
 ...





Model ::= *Param*⁺; *Data*⁺; *DistStmt*⁺;

Param ::= Param<Type>(...)

Data ::= Data<Type>(...)

DistStmt ::= Var /= DistExpr / Var /= Expr / Var /= Bexpr? Expr: Expr / observe(BExpr) / for (i=c1; i < c2; i++) { Var[i] = DistExpr }</pre>

DistExpr ::= bernoulli(Expr) | uniform(Expr, Expr) | normal(Expr, Expr)

BExpr ::= BExpr Boolop BExpr / ExprRelop Expr/ true / false

Expr ::= Expr ArithOp Expr / Var / c

Type ::= int | real | fixed < c, c> | vector < Type>

 $ArithOp \in \{+, -, *, /, **, ...\}, Boolop \in \{//, \&\&, ...\}, Relop \in \{<, ==, <=, ...\}$



Model ::= *Param*⁺; *Data*⁺; *DistStmt*⁺;

Param ::= Param < Type > (...)

Data ::= Data<Type>(...)

DistStmt ::= Var /= DistExpr / Var /= Expr / Var /= Bexpr? Expr: Expr / observe(BExpr) / for (i=c1; i < c2; i++) { Var[i] = DistExpr }</pre>

DistExpr ::= bernoulli(*Expr*) | uniform(*Expr*, *Expr*) | normal(*Expr*, *Expr*)

BExpr ::= BExpr Boolop BExpr / ExprRelop Expr/ true / false

Expr ::= Expr ArithOp Expr / Var / c

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 $ArithOp \in \{+, -, *, /, **, ...\}, Boolop \in \{//, \&\&, ...\}, Relop \in \{<, ==, <=, ...\}$



Model ::= *Param*⁺; *Data*⁺; *DistStmt*⁺;

Param ::= Param<Type>(...)

Data ::= Data<Type>(...)

DistStmt ::= Var /= DistExpr / Var /= Expr / Var /= Bexpr? Expr: Expr / observe(BExpr) / for (i=c1; i < c2; i++) { Var[i] = DistExpr }</pre>

DistExpr ::= bernoulli(Expr) | uniform(Expr, Expr) | normal(Expr, Expr)

BExpr ::= BExpr Boolop BExpr / ExprRelop Expr/ true / false

Expr ::= Expr ArithOp Expr / Var / c

Type ::= int | real | fixed < c, c> | vector < Type >

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Step 1) Fixed-Point Size Selection



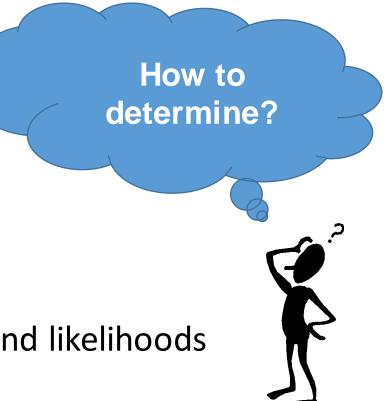
Fixed-Point Size Selection

- Fixed Point numbers given as <I,F>
- I is the amount of integer bits
- F is the amount of fractional bits
- Need enough integer bits for distributions and likelihoods



Fixed-Point Size Selection

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Fixed-Point Size Selection

- Fixed Point numbers given as <I,F>
- I is the amount of integer bits
- F is the amount of fractional bits



Use Interval

Analysis!



```
Param<Real> m;
Param<Real> b;
Data<Real> X = {1.2,2.4}
Data<Real> Y = {14.3,20.1}
```

```
m |= Uniform(0,10);
b |= Uniform(0,10);
Y |= Normal(m*X+b,0.25)
```



Param<Real> m; Param<Real> b; Data<Real> X = $\{1.2, 2.4\}$ X $\in [1.2, 2.4]$ Data<Real> Y = $\{14.3, 20.1\}$

```
m |= Uniform(0,10);
b |= Uniform(0,10);
Y |= Normal(m*X+b,0.25)
```



Param<Real> m; Param<Real> b; Data<Real> X = $\{1.2, 2.4\}$ Data<Real> Y = $\{14.3, 20.1\}$ Y $\in [14.3, 20.1]$

```
m |= Uniform(0,10);
b |= Uniform(0,10);
Y |= Normal(m*X+b,0.25)
```



Param<Real> m; Param<Real> b; Data<Real> X = $\{1.2, 2.4\}$ Data<Real> Y = $\{14.3, 20.1\}$ m |= Uniform(0,10); $m \in [0, 10]$ b |= Uniform(0,10); Y |= Normal(m*X+b, 0.25)



```
Param<Real> m;
Param<Real> b;
Data<Real> X = {1.2,2.4}
Data<Real> Y = {14.3,20.1}
```

m |= Uniform(0,10);
b |= Uniform(0,10);
Y |= Normal(
$$m \times X + b$$
, 0.25) $b \in [0, 10]$



```
Param<Real> m;

Param<Real> b;

Data<Real> X = {1.2,2.4}

Data<Real> Y = {14.3,20.1}

m |= Uniform(0,10);

b |= Uniform(0,10);

Y |= Normal(m*X+b,0.25) Y \in [-\infty, \infty]
```



Param<Real> m; Param<Real> b; Data<Real> X = {1.2,2.4} Data<Real> Y = {14.3,20.1} Y \in [14.3,20.1] m |= Uniform(0,10); b |= Uniform(0,10); Y |= Normal(m*X+b,0.25) Y \in [- ∞ , ∞]



 $Y \in [-\infty, \infty]$

Param<Real> m; Param<Real> b; Data<Real> X = $\{1.2, 2.4\}$ Data<Real> Y = $\{14.3, 20.1\}$ Y $\in [14.3, 20.1]$ m |= Uniform(0,10); b |= Uniform(0,10);

Need to take bigger of the two



```
Param<Real> m;
Param<Real> b;
Data<Real> X = {1.2,2.4}
Data<Real> Y = {14.3,20.1}
```

$$\begin{array}{ll} m & | = \text{Uniform}(0, 10); \\ b & | = \text{Uniform}(0, 10); \\ Y & | = \text{Normal}(m^*X + b, 0.25) \end{array} \quad \begin{array}{ll} \text{Nee} \\ Y \in [-\infty, \infty] \end{array} \quad \begin{array}{ll} C \\ C \end{array}$$

Need to take bigger of the two

Can we do better?



```
Param<Real> m;
Param<Real> b;
Data<Real> X = {1.2,2.4}
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```

m |= Uniform(0,10);
b |= Uniform(0,10);
Y |= Normal(m*X+b,0.25)
$$Y \in [-\infty, \infty]$$

Need to take bigger of the two

Can we do better?



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Param<Real> m;
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Data<Real> X = {1.2,2.4}
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```
 \begin{array}{l} m \mid = \text{Uniform}(0,10); \\ b \mid = \text{Uniform}(0,10); \\ Y \mid = \text{Normal}(m^*X+b,0.25) \end{array} \quad Y \in [-\infty,\infty] \end{array}
```

Samplers truncate to +/- 6σ



```
Param<Real> m;
Param<Real> b;
Data<Real> X = {1.2,2.4}
Data<Real> Y = {14.3,20.1}
```

```
m |= Uniform(0,10);
b |= Uniform(0,10);
Y |= Normal(m*X+b,0.25)
```

Samplers truncate to +/- 6o

 $Y \in [m^*X + b - 6^*0.25 \ , m^*X + b + 6^*0.25 \]$



```
Param<Real> m;

Param<Real> b;

Data<Real> X = {1.2,2.4}

Data<Real> Y = {14.3,20.1}

m |= Uniform(0,10);

b |= Uniform(0,10);

Y |= Normal(m*X+b,0.25) Y \in [m*X+b-6*0.25, m*X+b+6*0.25]
```

Simplify using Interval Arithmetic



```
Param<Real> m;
Param<Real> b;
Data<Real> X = {1.2,2.4}
Data<Real> Y = {14.3,20.1}
m |= Uniform(0,10);
b |= Uniform(0,10);
Y |= Normal(m*X+b,0.25) Y ∈ [-1.5,35.5]
```

Simplify using Interval Arithmetic



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Param<Real> m;
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```
m |= Uniform(0,10);
b |= Uniform(0,10);
Y |= Normal(m*X+b,0.25)
```

 $X \in [1.2, 2.4]$ $m \in [0, 10]$ $b \in [0, 10]$ $Y \in [-1.5, 35.5]$

Final Intervals



- MCMC requires computing log-likelihoods for acceptance ratio
- Each distribution has different likelihood
- Need Summation over all data samples



 $egin{aligned} \log{(acc)} &= \sum_{i=1}^m (\log{(\Pr(d_i \,|\, \mathbf{x_p}))} + \log{(\Pr(\mathbf{x_p}))} + \log{(q(\mathbf{x_c} \,|\, \mathbf{x_p}))}) \ &- \sum_{i=1}^m (\log{(\Pr(d_i \,|\, \mathbf{x_c}))} + \log{(\Pr(\mathbf{x_c}))} + \log{(q(\mathbf{x_p} \,|\, \mathbf{x_c}))}) \end{aligned}$



Acceptance Ratio

$$egin{aligned} \log\left(acc
ight) &= \sum_{i=1}^{m} (\log\left(\Pr(d_i \,|\, \mathbf{x_p})
ight) + \log\left(\Pr(\mathbf{x_p})
ight) + \log\left(q(\mathbf{x_c} \,|\, \mathbf{x_p})
ight)) \ &- \sum_{i=1}^{m} (\log\left(\Pr(d_i \,|\, \mathbf{x_c})
ight) + \log\left(\Pr(\mathbf{x_c})
ight) + \log\left(q(\mathbf{x_p} \,|\, \mathbf{x_c})
ight)) \end{aligned}$$



Log-likelihoods

$$\begin{split} \log{(acc)} &= \sum_{i=1}^{m} \left(\log{(\Pr(d_i \mid \mathbf{x_p}))} + \log{(\Pr(\mathbf{x_p}))} + \log{(q(\mathbf{x_c} \mid \mathbf{x_p}))} \right) \\ &- \sum_{i=1}^{m} \left(\log{(\Pr(d_i \mid \mathbf{x_c}))} + \log{(\Pr(\mathbf{x_c}))} + \log{(q(\mathbf{x_p} \mid \mathbf{x_c}))} \right) \end{split}$$



Summation over **all** observed data

$$\log (acc) = \sum_{i=1}^{m} (\log (\Pr(d_i \mid \mathbf{x_p})) + \log (\Pr(\mathbf{x_p})) + \log (q(\mathbf{x_c} \mid \mathbf{x_p}))) \\ - \sum_{i=1}^{m} (\log (\Pr(d_i \mid \mathbf{x_c})) + \log (\Pr(\mathbf{x_c})) + \log (q(\mathbf{x_p} \mid \mathbf{x_c})))$$



• How to bound terms like $(\log (\Pr(d_i | \mathbf{x_c})) \text{ and } \log (\Pr(\mathbf{x_c}))$?



- How to bound terms like $(\log (\Pr(d_i | \mathbf{x_c})) \text{ and } \log (\Pr(\mathbf{x_c}))$?
- Propagate interval bounds through each distribution's likelihood



- How to bound terms like $(\log (\Pr(d_i | \mathbf{x_c})) \text{ and } \log (\Pr(\mathbf{x_c}))$?
- Propagate interval bounds through each distribution's likelihood
- Leverage previously computed intervals!





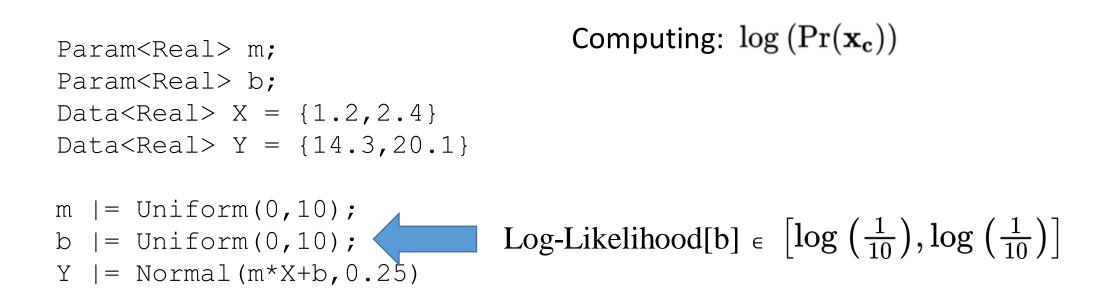
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Param<Real> b;
Data<Real> X = {1.2,2.4}
Data<Real> Y = {14.3,20.1}
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m |= Uniform(0,10);
b |= Uniform(0,10);
Y |= Normal(m*X+b,0.25)
```



Param<Real> m; Param<Real> b; Data<Real> X = {1.2,2.4} Data<Real> Y = {14.3,20.1} m |= Uniform(0,10); b |= Uniform(0,10); Y |= Normal(m*X+b,0.25) Computing: log ($Pr(x_c)$) Log-Likelihood[m] $\in \left[log(\frac{1}{10}), log(\frac{1}{10})\right]$







```
Param<Real> m;

Param<Real> b;

Data<Real> X = \{1.2, 2.4\}

Data<Real> Y = \{14.3, 20.1\}

m |= Uniform(0,10);

b |= Uniform(0,10);

Y |= Normal(m*X+b,0.25)

Computing: (log(Pr(d_i | \mathbf{x}_c))

Computing: (log(Pr(d_i | \mathbf{x}_c))

Log-Likelihood[Y] \in ???
```



```
Param<Real> m;
Param<Real> b;
Data<Real> X = {1.2,2.4}
Data<Real> Y = {14.3,20.1}
m |= Uniform(0,10);
b |= Uniform(0,10);
Y |= Normal(m*X+b,0.25)
Leverage previously computed intervals!
```



```
Param<Real> m;
Param<Real> b;
Data<Real> X = {1.2,2.4}
Data<Real> Y = {14.3,20.1}
```

 $X \in [1.2, 2.4]$ $m \in [0, 10]$ $b \in [0, 10]$ $Y \in [-1.5, 35.5]$

```
m |= Uniform(0,10);
b |= Uniform(0,10);
Y |= Normal(m*X+b,0.25)
```

Log-Likelihood[Y] \in



```
Param<Real> m;

Param<Real> b;

Data<Real> X = {1.2,2.4}

Data<Real> Y = {14.3,20.1}

M = Uniform(0,10);

b = Uniform(0,10);

Y = Normal(m*X+b,0.25)

X \in [1.2, 2.4]

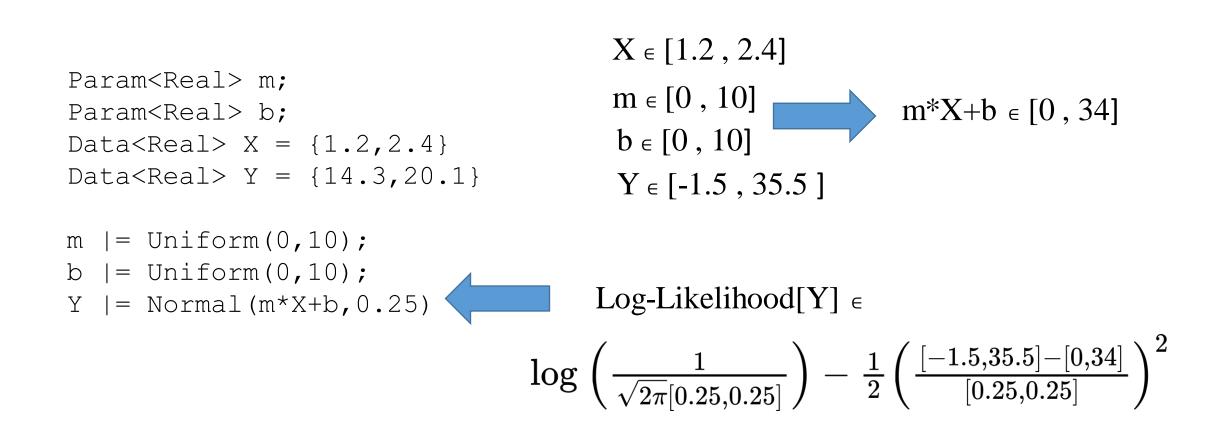
m \in [0, 10]

b \in [0, 10]

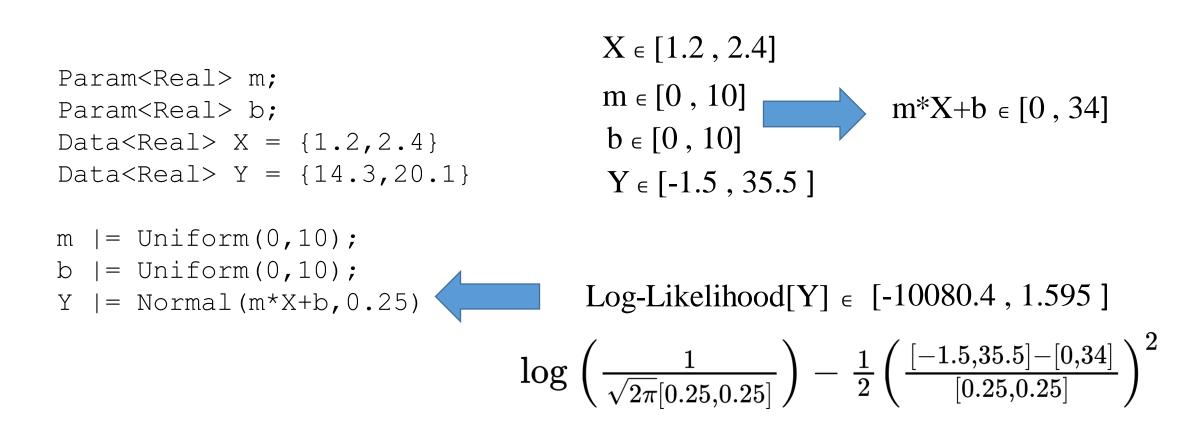
Y \in [-1.5, 35.5]

M = Log-Likelihood[Y] \in Comparison of the second state of th
```











• Still need to bound proposal kernel terms:

 $\log\left(q(\mathbf{x}_{\mathbf{p}} \mid \mathbf{x}_{\mathbf{c}})\right)) \quad \& \quad \log\left(q(\mathbf{x}_{\mathbf{c}} \mid \mathbf{x}_{\mathbf{p}})\right)$



• Still need to bound proposal kernel terms:

 $\log\left(q(\mathbf{x}_{\mathbf{p}} \,|\, \mathbf{x}_{\mathbf{c}})\right)) \quad \& \quad \log\left(q(\mathbf{x}_{\mathbf{c}} \,|\, \mathbf{x}_{\mathbf{p}})\right)$

- Proposal kernel has known form (Normal, uniform, etc.)
- $\mathbf{x}_{\mathbf{p}}$ and $\mathbf{x}_{\mathbf{c}}$ have known non-infinite interval bounds
- Proposal is **symmetric**



Computing: $\log (q(\mathbf{x}_{\mathbf{p}} | \mathbf{x}_{\mathbf{c}})))$

```
Param<Real> m;
Param<Real> b;
Data<Real> X = {1.2,2.4}
Data<Real> Y = {14.3,20.1}
```

```
m |= Uniform(0,10);
b |= Uniform(0,10);
Y |= Normal(m*X+b,0.25)
```

When proposal kernel is Normal(**x**_c,1)



Computing: $\log (q(\mathbf{x}_{\mathbf{p}} | \mathbf{x}_{\mathbf{c}})))$

```
Param<Real> m;
Param<Real> b;
Data<Real> X = {1.2,2.4}
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```
m |= Uniform(0,10);
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When proposal kernel is Normal(**x**_c,1)

> $m \in [0, 10]$ $b \in [0, 10]$



Computing: $\log (q(\mathbf{x}_{\mathbf{p}} | \mathbf{x}_{\mathbf{c}})))$

```
Param<Real> m;
Param<Real> b;
Data<Real> X = {1.2,2.4}
Data<Real> Y = {14.3,20.1}
```

When proposal kernel is Normal(**x**_c,1)

> $m \in [0, 10]$ $b \in [0, 10]$

m |= Uniform(0,10); b |= Uniform(0,10); Y |= Normal(m*X+b,0.25)

 $Log-Likelihood[m] = Log-Likelihood[b] \in$

$$\log\left(rac{1}{\sqrt{2\pi}[1,1]}
ight) - rac{1}{2} igg(rac{[0,10]-[0,10]}{[1,1]}igg)^2$$



Computing: $\log (q(\mathbf{x}_{\mathbf{p}} | \mathbf{x}_{\mathbf{c}})))$ When proposal kernel is $Normal(\mathbf{x}_{c}, 1)$ $m \in [0, 10]$ $b \in [0, 10]$ Log-Likelihood[m]= Log-Likelihood[b] ∈ $\log\left(rac{1}{\sqrt{2\pi}[1,1]}
ight) - rac{1}{2}\left(rac{[0,10]-[0,10]}{[1,1]}
ight)^2$

Param<Real> m;
Param<Real> b;
Data<Real> X = {1.2,2.4}
Data<Real> Y = {14.3,20.1}

```
m |= Uniform(0,10);
b |= Uniform(0,10);
Y |= Normal(m*X+b,0.25)
```



Computing: $\log (q(\mathbf{x}_{\mathbf{p}} | \mathbf{x}_{\mathbf{c}})))$

```
Param<Real> m;
Param<Real> b;
Data<Real> X = {1.2,2.4}
Data<Real> Y = {14.3,20.1}
```

When proposal kernel is Normal(\mathbf{x}_{c} , 1) $m \in [0, 10]$

b ∈ [0, 10]

m |= Uniform(0,10); b |= Uniform(0,10); Y |= Normal(m*X+b,0.25)

Log-Likelihood[m]= Log-Likelihood[b] \in [-50.9, -0.91]

$$\log\left(rac{1}{\sqrt{2\pi}[1,1]}
ight) - rac{1}{2} igg(rac{[0,10]-[0,10]}{[1,1]}igg)^2$$



• We can now bound all terms for

 $egin{aligned} \log{(acc)} &= \sum_{i=1}^m (\log{(\Pr(d_i \,|\, \mathbf{x_p}))} + \log{(\Pr(\mathbf{x_p}))} + \log{(q(\mathbf{x_c} \,|\, \mathbf{x_p}))}) \ &- \sum_{i=1}^m (\log{(\Pr(d_i \,|\, \mathbf{x_c}))} + \log{(\Pr(\mathbf{x_c}))} + \log{(q(\mathbf{x_p} \,|\, \mathbf{x_c}))}) \end{aligned}$



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 $\left[-10080.4, 1.595
ight]$



• We can now bound all terms for

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$$\left[\log\left(\frac{1}{10}\right), \log\left(\frac{1}{10}\right)\right] + \left[\log\left(\frac{1}{10}\right), \log\left(\frac{1}{10}\right)\right]$$



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[-50.9, -0.91]



• We can now bound all terms for

$$egin{aligned} \log\left(acc
ight) &= \sum_{i=1}^{m} \left(\log\left(\Pr(d_i \,|\, \mathbf{x_p})
ight) + \log\left(\Pr(\mathbf{x_p})
ight) + \log\left(q(\mathbf{x_c} \,|\, \mathbf{x_p})
ight)
ight) \ &- \sum_{i=1}^{m} \left(\log\left(\Pr(d_i \,|\, \mathbf{x_c})
ight) + \log\left(\Pr(\mathbf{x_c})
ight) + \log\left(q(\mathbf{x_p} \,|\, \mathbf{x_c})
ight)
ight) \end{aligned}$$

$$\left[-10135.50, -3.92
ight]$$



• We can now bound all terms for

$$egin{aligned} \log\left(acc
ight) &= \sum_{i=1}^{m} (\log\left(\Pr(d_i \,|\, \mathbf{x_p})
ight) + \log\left(\Pr(\mathbf{x_p})
ight) + \log\left(q(\mathbf{x_c} \,|\, \mathbf{x_p})
ight)) \ &- \sum_{i=1}^{m} (\log\left(\Pr(d_i \,|\, \mathbf{x_c})
ight) + \log\left(\Pr(\mathbf{x_c})
ight) + \log\left(q(\mathbf{x_p} \,|\, \mathbf{x_c})
ight)) \end{aligned}$$

[-10135.50, -3.92] * Number of Observations



Can we get away with less?



Can we get away with less?

Yes!



Can we get away with less?

Yes! Overflows are ok.



• Fixed-Point uses **2's complement** integer arithmetic



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- Wrap-around overflows in Likelihood Summation are **ok...**



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...provided **final** result is within representable range

Only need enough integer bits for largest single likelihood



• Given interval bounds how do we choose final size?



- Given interval bounds how do we choose final size?
- Need a different size for distributions and likelihoods!



- Given interval bounds how do we choose final size?
- Need a different size for <u>distributions</u> and likelihoods!

$$I_M \geq \log_2{(\left\lceil\max_{v\in Vars}(|\underline{v}|,|\overline{v}|)
ight
ceil)}$$
 Integer Bits



- Given interval bounds how do we choose final size?
- Need a different size for <u>distributions</u> and likelihoods!

$$egin{aligned} &I_M \geq \log_2{(\left\lceil\max_{v\in Vars}(|ec{v}|,|ec{v}|)
ight
ceil)} \ &F_M = 32 - 1 - I_M \end{aligned}$$
 Fractional Bits



- Given interval bounds how do we choose final size?
- Need a different size for distributions and <u>likelihoods</u>!

 $I_{LL} \geq \log_2\left(\left\lceil\max_{v\in Vars}(| ext{log-likelihood}(\underline{v})|, | ext{log-likelihood}(\overline{v})|)
ight
ceil
ight)$



- Given interval bounds how do we choose final size?
- Need a different size for distributions and <u>likelihoods</u>!

 $I_{LL} \geq \log_2{(\left\lceil\max_{v\in Vars}(| ext{log-likelihood}(ec{v})|, | ext{log-likelihood}(\overline{v})|)
ight
ceil)})$

$$F_{LL} = 32 - 1 - I_{LL}$$



Step 2) MCMC Code Instrumentation & Optimization



Dynamic Checks

• Overflows need to be checked for at runtime



Dynamic Checks

- Overflows need to be checked for at runtime
- Luckily, we don't have to check *every* arithmetic operation



Dynamic Checks

- Overflows need to be checked for at runtime
- Luckily, we don't have to check *every* arithmetic operation
- Only check *final* Acceptance Ratio summation



Domain Specific Optimizations

• MCMC sampling benefits from:



Domain Specific Optimizations

- MCMC sampling benefits from:
- Constant Propagation through the Bayesian Network



Domain Specific Optimizations

- MCMC sampling benefits from:
- Constant Propagation through the Bayesian Network
- Memoization during likelihood computation



How does Statheros perform?



Evaluation - Methodology

- Took multiple benchmarks from the Literature
- Run MCMC for 10K samples + 5k burn-in to get posterior

Accuracy:	<u>True Param Value - Posterior Mean</u>
	True Param Value



Evaluation - Methodology

- Measure inference runtime and accuracy for Fixed Point (Statheros) against Float (32-bit) and Double (64-bit)
- Evaluated on 3 devices: Arduino (no FPU), Raspberry Pi and PocketBeagle









Evaluation - Methodology

- Measure inference runtime and accuracy for Fixed Point (Statheros) against Float (32-bit) and Double (64-bit)
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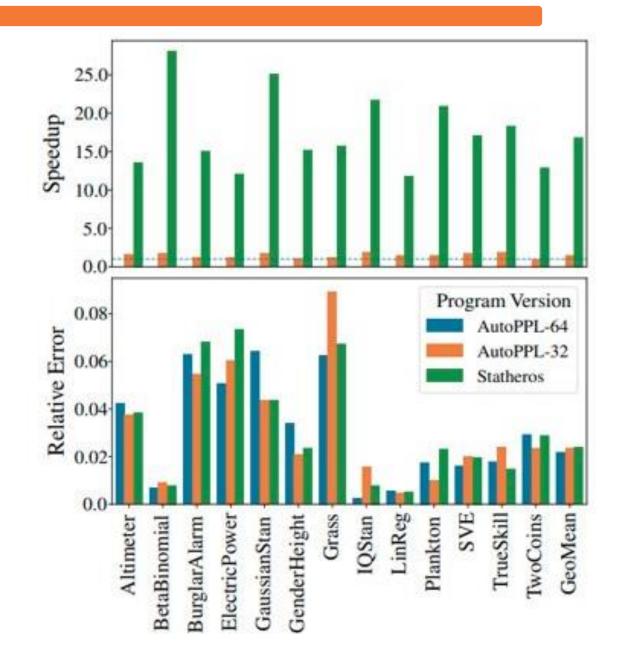


Benchmark	Distributions		
Altimeter	Bernoulli		
Beta-Binomial	Beta, Binomial		
Burglar Alarm	Bernoulli		
Electric Power	Bernoulli		
Gaussian Stan	Gaussian		
Gender Height	Bernoulli, Gaussian		
Grass	Bernoulli		
IQ Stan	Uniform, Gaussian		
Linear Regression	Uniform, Gaussian		
Plankton	Uniform, Gaussian		
SVE	Uniform, Triangular, Gaussian		
TrueSkill	Bernoulli, Gaussian		
TwoCoins	Bernoulli		



Arduino

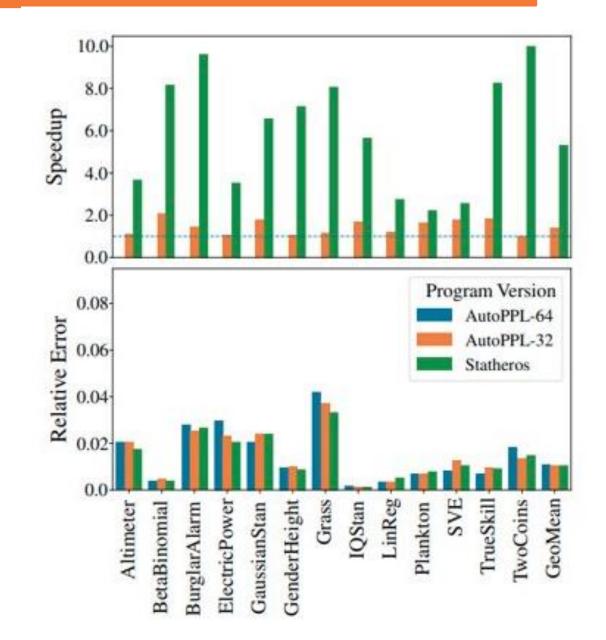
- Substantial speedup due to Arduino's lack of FPU
- Statheros GeoMean Speedup: 16.91x (over 64 bit double) 11.54x (over 32 bit float)
- Geomean Relative Error Statheros: 0.0239
 32 bit float: 0.0238
 64 bit float: 0.0218





PocketBeagle

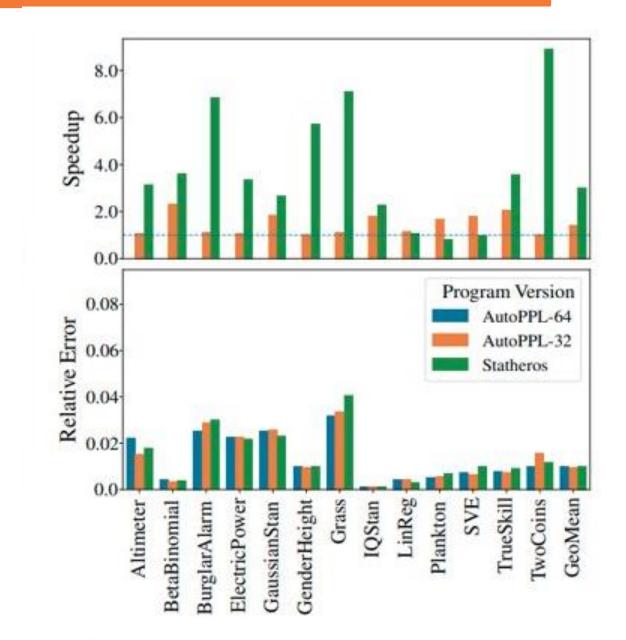
- PocketBeagle has low-end FPU -> still a large speedup
- Statheros GeoMean Speedup: 5.33x (over 64 bit double) 3.77x (over 32 bit float)
- Geomean Relative Error
 Statheros: 0.01
 32 bit float: 0.01
 64 bit float: 0.01





Raspberry Pi

- Raspberry Pi does have an FPU: speedup not as large
- Statheros GeoMean Speedup: 3.04x (over 64 bit double) 2.15x (over 32 bit float)
- Geomean Relative Error
 Statheros: 0.01
 32 bit float: 0.01
 64 bit float: 0.01



Benchmark	Parameter Configuration		Likelihood Configuration	
	Integer	Fractional	Integer	Fractional
Altimeter	7	24	7	24
Beta-Binomial	7	24	19	12
Burglar Alarm	7	24	7	24
Electric Power	7	24	7	24
Gaussian Stan	11	20	19	12
Gender Height	11	20	11	20
Grass	7	24	7	24
IQ Stan	11	20	19	12
Linear Regression	7	24	19	12
Plankton	7	24	7	24
SVE	7	24	19	12
TrueSkill	11	20	7	24
TwoCoins	7	24	7	24



Evaluation - Takeaways

- Statheros faster than float and double on all benchmarks for Arduino and PocketBeagle
- Inferred Fixed Point configurations tolerate approximation
- Not all overflows are bad!



More in the Paper:

- Detailed Algorithmic Description of Compilation
- Impact of Optimizations
- Discussion of Related Work



Statheros Takeaways

• Probabilistic Programming + Fixed Point Precision offers



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Major Runtime Savings!



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